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# The Detection of a Sex Difference in Recombination Values Using Double Heterozygotes 

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#### Abstract

The method of detecting or measuring a difference between the recombination fractions in female and male gametogenesis is found on inserting typical numerical values to be quite as insensitive as its critics have suggested. If, how'ever, double heterozygotes in Coupling and Repulsion are both available, reciprocal intercrosses will determine both recombination fractions without ambiguity. The use of double heterozygotes for studying linkage has thus fewer disadvantages, when both kinds are available, than has been supposed.


## r. Linkage Values obtained by Self-fertilization

It has been observed by many authors (Fisher \& Balmukand, 1928) that when a recombination fraction is determined by selfing double heterozygotes, what is calculable from such data is not the actual recombination value in male or female gametogenesis, unless these are equal, but the geometric mean of these two values. More rarely the comment has been added that if such data are available, not only from heterozygotes in Coupling but also from those in Repulsion, there will be a discrepancy between the linkage values obtained from these two sources, so that, in theory at least, a difference between the female and the male values could be detected in this way.

It might be guessed that it would require enormous numbers to detect, and still more to evaluate with any precision, a sex difference in this way. It was early observed also that even if a difference were recognizable, such data would supply no indication whether the recombination fractions were higher in the female or in the male.

Since I believe no case in the literature has been exhibited in detail, and since a general algebraic treatment would be complex and barely profitable, it may be worth the while to set out the numerical calculations for a sufficiently typical example, so that the provisional conclusions above may be seen to be verifiable and the steps of such a verification made clear.

The genetic situation may be expressed in a short table of gametic frequencies:

|  | $q$ |  | gametogenesis |  |
| :--- | :---: | :---: | :---: | :---: |
|  | Coupling | Repulsion | O Coupling $^{\prime}$ gamenesis |  |
|  | Repulsion |  |  |  |
| $A B$ | $q / 2$ | $p / 2$ | $q^{\prime} / 2$ | $p^{\prime} / 2$ |
| $A b$ | $p / 2$ | $q / 2$ | $p^{\prime} / 2$ | $q^{\prime} / 2$ |
| $a B$ | $p / 2$ | $q / 2$ | $p^{\prime} / 2$ | $q^{\prime} / 2$ |
| $a b$ | $q / 2$ | $p / 2$ | $q^{\prime} / 2$ | $p^{\prime} / 2$. |

If dominance is complete in both factors, the phenotypic frequencies arising from self-fertilization are:

|  | Coupling | Repulsion |
| :--- | :---: | :---: |
| $A B$ | $\left(2+q q^{\prime}\right) / 4$ | $\left(2+p p^{\prime}\right) / 4$ |
| $A b$ | $\left(\mathrm{I}-q q^{\prime}\right) / 4$ | $\left(\mathrm{i}-p p^{\prime}\right) / 4$ |
| $a B$ | $\left(\mathrm{r}-q q^{\prime}\right) / 4$ | $\left(\mathrm{~s}-p p^{\prime}\right) / 4$ |
| $a b$ | $q q^{\prime} / 4$ | $p p^{\prime} / 4$. |

In coupling, the frequencies expected in all observable classes are expressible in terms of $q q^{\prime}$, which may, therefore, be estimated from such data: in repulsion all we can estimate is $p p^{\prime}$. Now

$$
p+q=\mathbf{1}=p^{\prime}+q^{\prime},
$$

so that if male and female gametogenesis yield the same gametic ratios, we shall find

$$
\sqrt{q q^{\prime}}+\sqrt{p p^{\prime}}=\mathrm{I}
$$

In general, however, there will be, quite apart from errors of random sampling, a discrepancy in the expectations measured by

$$
\mathrm{I}-\sqrt{q q^{\prime}}-\sqrt{p p^{\prime}}
$$

We may calculate the conditions in which such a discrepancy between the appropriate estimates will be significant.

If, for example, the recombination fraction were $30 \%$ in females, but $20 \%$ in males we should have

$$
\begin{aligned}
& \sqrt{p p^{\prime}}=\sqrt{\cdot 06}=\cdot 24494,897 \\
& \sqrt{q q^{\prime}}=\sqrt{\cdot 56}=74833,148 \\
& \text { Discrepancy }=-00671,955 \\
& \text { Total } \quad \underset{1 \cdot 00000,000 .}{ }
\end{aligned}
$$

At the $5 \%$ level of significance the discrepancy exceeds its standard error in the ratio

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| :--- |
| $A B$ |

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| $A b$ | $\left(\mathrm{I}-q q^{\prime}\right) / 4$ | $\left(\mathrm{x}-p p^{\prime}\right) / 4$ |
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At the $5 \%$ level of significance the discrepancy exceeds its standard error in the ratio

For a reasonable chance of significance it is necessary that the standard error of the estimated value of

$$
\sqrt{p p^{\prime}}+\sqrt{q q^{\prime}}
$$

shall be brought down to approximately

$$
\frac{.0067 \mathrm{I}, 955}{\mathrm{I} \cdot 9599^{6}}=\cdot 00342,84 \mathrm{I}
$$

Now if $\theta$ stand for $p p^{\prime}$ or $q q^{\prime}$, the sampling variance of $\sqrt{\theta}$ as estimated is known (Fisher, 1928-1958) to be

$$
V(\sqrt{\theta})=\frac{(\mathrm{I}-\theta)(2+\theta)}{2 n(\mathrm{I}+2 \theta)}
$$

where the estimate is based on $n$ observations.
For values of $\theta, \cdot 06$ and $\cdot 56$, the coefficients of $\mathrm{I} / n$ are

$$
\begin{array}{cl}
\theta & n V(\sqrt{ } \bar{\theta}) \\
.06 & \cdot 86446,429 \\
.56 & \cdot 26566,038 .
\end{array}
$$

The variance of the sum is, therefore,

$$
\frac{\mathrm{I}}{n_{1}}(\cdot 86446,429)+\frac{\mathrm{I}}{n_{2}}(\cdot 26566,038),
$$

where $n_{1}$ offspring are bred from parents in Repulsion, and $n_{2}$ from parents in Coupling. To minimize this for a given total $N$ of the two groups, take

$$
\frac{n_{1}}{\sqrt{\cdot 8644^{6,429}}}=\frac{n_{2}}{\sqrt{\cdot 26566,038}}
$$

or

$$
\frac{n_{1}}{92976,57^{2}}=\frac{n_{2}}{55154^{2,253}}=\frac{N}{1 \cdot 44518,825}
$$

and when the optimal allocation has been made, the Standard Error will be

$$
\frac{1.44518,825}{\sqrt{\bar{N}}}
$$

and the necessary value of $N$

$$
\left(\frac{\mathrm{x} \cdot 44518,825}{\cdot 0034284^{\mathrm{x}}}\right)^{2}=(421 \cdot 533)^{2}=177,690
$$

confirming that the method is insensitive to the point of impracticability.

## 2. The Use of Intercrosses

It seems not to have been noticed that with no more experimental material than was postulated above-both types of double heterozygote, but no possibility of back-crossing, for the double recessives may be sterile-it is still possible to detect and to measure a sex difference in the recombination fraction, by making reciprocal intercrosses between the two types of double heterozygotes. Moreover by this path, when a difference exists there is no ambiguity of interpretation: it is clear which sex shows closer linkage. The reciprocal crosses will yield the phenotypic frequencies shown below.

|  | Coupling ovules <br> $\times$ <br> repulsion pollen | Repulsion ovules <br> $\times$ coupling pollen |
| :--- | :---: | :---: |
| $A B$ | $\left(2+p^{\prime} q\right) / 4$ | $\left(2+p q^{\prime}\right) / 4$ |
| $A b$ | $\left(\mathrm{I}-p^{\prime} q\right) / 4$ | $\left(\mathrm{I}-p q^{\prime}\right) / 4$ |
| $a B$ | $\left(\mathrm{r}-p^{\prime} q\right) / 4$ | $\left(\mathrm{I}-p q^{\prime}\right) / \dot{4}$ |
| $a b$ | $p^{\prime} q / 4$ | $p q^{\prime} / 4$. |

Using the recombinations chosen above,

$$
p^{\prime} q=\cdot 24, p q^{\prime}=\cdot 14
$$

the difference is

$$
\cdot 10
$$

and for significance this needs a standard error not greater than

$$
\cdot 0510214
$$

The sampling variance of either of the estimates is

$$
2 \theta(\mathrm{r}-\theta)(2+\theta)+n(\mathrm{r}+2 \theta)
$$

which gives, for the values $\cdot 24$ and $\cdot 14$, the coefficients of $\mathrm{I} / n_{1}$ and $\mathrm{I} / n_{2}$

$$
\cdot 55212,973 ; \cdot 40258,750
$$

of which the square roots are

$$
\cdot 74305,433 ; \cdot 63449,783
$$

giving the relative numbers to be bred from the two crosses to obtain optimal precision.

The sum is

$$
\mathrm{x} \cdot 37755,216
$$

and this divided by

$$
0.0510214
$$

gives

$$
\begin{gathered}
\sqrt{N} \quad 26.9995 \\
N 7_{228.973}
\end{gathered}
$$

so that only a few hundred from each cross would serve to detect the sex differences, and a few thousand plants would give separate estimates of the recombination fraction of female and male gametogenesis.

In practice the scoring coefficients given in "Statistical Tables" (Fisher \& Yates, 1957), Table XIII could be used to evaluate $\sqrt{p^{\prime} q}$ and $\sqrt{p q^{\prime}}$.

## REFERENCES

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